

VE401 RECITATION CLASS NOTE9

Single Sample Test

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1 Parametric Test for Mean

Distribution of X_i	Sample size n	Variance σ^2	Statistic
$X_i \sim \mathcal{N}(\mu, \sigma)$	any	known	$\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \sim \mathcal{N}(0, 1)$
$X_i \sim$ any distribution	large	known	$\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \sim \mathcal{N}(0, 1)$
$X_i \sim$ any distribution	large	unknown	$\frac{\bar{X} - \mu}{\frac{s}{\sqrt{n}}} \sim \mathcal{N}(0, 1)$
$X_i \sim \mathcal{N}(\mu, \sigma)$	small	unknown	$\frac{\bar{X} - \mu}{\frac{s}{\sqrt{n}}} \sim t_{n-1}$
$X_i \sim$ any distribution	small	known or unknown	Go home!

1.1 Z-test

Test Statistic:

$$Z = \frac{\bar{X} - \mu_0}{\sigma / \sqrt{n}}$$

Reject at significance level α :

- (i) $H_0 : \mu = \mu_0$ if $|Z| > z_{\alpha/2}$
- (ii) $H_0 : \mu \leq \mu_0$ if $Z > z_{\alpha}$
- (iii) $H_0 : \mu \geq \mu_0$ if $Z < z_{-\alpha}$

Abscissa of OC Curves:

$$d = \frac{|\mu - \mu_0|}{\sigma}$$

1.2 T-test

Test Statistic:

$$T_{n-1} = \frac{\bar{X} - \mu_0}{S/\sqrt{n}}$$

Reject at significance level α :

- (i) $H_0 : \mu = \mu_0$ if $|T_{n-1}| > t_{\alpha/2}$
- (ii) $H_0 : \mu \leq \mu_0$ if $T_{n-1} > t_{\alpha}$
- (iii) $H_0 : \mu \geq \mu_0$ if $T_{n-1} < t_{-\alpha}$

Abscissa of OC Curves:

$$d = \frac{|\mu - \mu_0|}{\sigma}$$

Estimate σ :

- (i) Use prior experiments to insert a rough estimate for σ
- (ii) Express the difference $\delta = |\mu - \mu_0|$ relative to σ
- (iii) Substitute the sample standard deviation s for σ .

2 Parametric Test for Variance

Distribution of X_i	Sample size n	Mean μ	Statistic
$X_i \sim \mathcal{N}(\mu, \sigma)$	any	known or unknown	$\frac{s^2(n-1)}{\sigma^2} \sim \chi_{n-1}^2$

2.1 Chi-squared Test

Test Statistic:

$$\chi_{n-1}^2 = \frac{(n-1)S^2}{\sigma_0^2}$$

Reject at significance level α :

- (i) $H_0 : \sigma = \sigma_0$ if $\chi_{n-1}^2 > \chi_{\alpha/2, n-1}^2$ or $\chi_{n-1}^2 < \chi_{1-\alpha/2, n-1}^2$
- (ii) $H_0 : \sigma \leq \sigma_0$ if $\chi_{n-1}^2 > \chi_{\alpha, n-1}^2$
- (iii) $H_0 : \sigma \geq \sigma_0$ if $\chi_{n-1}^2 < \chi_{1-\alpha, n-1}^2$

Abscissa of OC Curves:

$$d = \frac{\sigma}{\sigma_0}$$

3 Non-parametric Test for Median

1. Non-parametric statistics: do not assume the dependence on any parameter.
2. Distribution-free statistics: do not assume that X follows any particular distribution (such as the normal distribution).

3.1 Sign Test

Number of Signs:

Let X_1, \dots, X_n be a random sample of size n from an arbitrary continuous distribution and let

$$Q_+ = \#\{X_k : X_k - M_0 > 0\}, \quad Q_- = \#\{X_k : X_k - M_0 < 0\}$$

Q_+ is the number of “positive signs” and Q_- the number of “negative signs”.

If $X_i - M_0 = 0$, usual practice is to exclude X_i from the analysis.

Calculating P-value:

$$P[Q_- \leq k \mid M = M_0] = \sum_{x=0}^k \binom{n}{x} \frac{1}{2^n}$$

Rejecting H_0 :

We reject at significance level α :

- (i) $H_0 : M \leq M_0$ if $P[Q_- < k \mid M = M_0] < \alpha$
- (ii) $H_0 : M \geq M_0$ if $P[Q_+ < k \mid M = M_0] < \alpha$
- (iii) $H_0 : M = M_0$ if $P[\min(Q_-, Q_+) < k \mid M = M_0] < \alpha/2$

3.2 Wilcoxon Signed Rank Test

Sum of Ranks:

Let X_1, \dots, X_n be a random sample of size n from a symmetric distribution.

Order the n absolute differences $|X_i - M_0|$ according to magnitude, so that $X_{R_i} - M_0$ is the R_i th smallest difference by modulus.

If ties in the rank occur, the mean of the ranks is assigned to all equal values.

Let

$$W_+ = \sum_{R_i > 0} R_i, \quad |W_-| = \sum_{R_i < 0} |R_i|$$

W_+ is the sum of “positive ranks” and W the modulus of the sum of “negative signs”.

If $X_i - M_0 = 0$, usual practice is to exclude X_i from the analysis.

The Critical Value:

For small n , for example ≤ 30 , critical values are given in tables.

Rejecting H_0 :

We reject at significance level α :

- (i) $H_0 : M \leq M_0$ if W_- is smaller than the critical value for α
- (ii) $H_0 : M \geq M_0$ if W_+ is smaller than the critical value for α
- (iii) $H_0 : M = M_0$ if $W = \min(W_+, |W_-|)$ is smaller than the critical value for $\alpha/2$

Calculating P-Value:

For large n , we use the normal distribution to approximately calculate P-value, with parameters

$$\mu = \frac{n(n+1)}{4}, \quad \sigma^2 = \frac{n(n+1)(2n+1)}{24}$$

Rejecting H_0 :

We reject at the [P-value] significance level if [P-value] is small.

4 Inferences on Proportion

4.1 Estimator and Statistic

Define the random variable:

$$X = \begin{cases} 1 & \text{has trait,} \\ 0 & \text{does not have trait.} \end{cases}$$

If we take a random sample X_1, \dots, X_n of X , the sample mean is an (unbiased) estimator for p :

$$\hat{p} = \bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

By the central limit theorem, when n is large enough, \hat{p} is approximately normally distributed with mean p and variance $p(1 - p)/n$. Hence,

$$Z = \frac{\hat{p} - p}{\sqrt{p(1 - p)/n}}$$

4.2 Interval Estimation and Sample Size

It follows immediately that the following is a $100(1 - \alpha)\%$ confidence interval for p :

$$\hat{p} \pm z_{\alpha/2} \sqrt{p(1 - p)/n}$$

But the interval depends on the unknown parameter p , which we are actually trying to estimate! One solution to the problem is to replace p by \hat{p} , i.e., to write

$$\hat{p} \pm z_{\alpha/2} \sqrt{\hat{p}(1 - \hat{p})/n}$$

We may want to be able to claim that “with $100(1 - \alpha)\%$ probability, \hat{p} differs from p by at most d .”

Given a $100(1 - \alpha)\%$ confidence interval $p = \hat{p} \pm z$ know with $100(1 - \alpha)\%$ confidence that

$$d = z_{\alpha/2} \sqrt{\hat{p}(1 - \hat{p})/n}$$

Given d , this means that we should choose

$$n = \frac{z_{\alpha/2}^2 \hat{p}(1 - \hat{p})}{d^2}$$

which requires us to have an estimate \hat{p} of p beforehand.

4.3 Test for Proportion

Let X_1, \dots, X_n be a random sample of (large) size n from a Bernoulli distribution with parameter p and let $\hat{p} = \bar{X}$ denote the sample mean. Then any test based on the statistic

$$Z = \frac{\hat{p} - p_0}{\sqrt{p_0(1 - p_0)/n}}$$

is called a large-sample test for proportion.

We reject at significance level α :

- (i) $H_0 : p = p_0$ if $|Z| > z_{\alpha/2}$
- (ii) $H_0 : p \leq p_0$ if $Z > z_\alpha$
- (iii) $H_0 : p \geq p_0$ if $Z < -z_\alpha$

5 Acceptance Sampling

Acceptance Sampling

A construction firm receives a shipment of $N = 50$ steel rods to be used in the construction of a bridge. The lot must be checked to ensure that the breaking strength of the rods meets specifications. The lot will be rejected if among the 200 rods more than 10% fail to meet specifications. We define the true proportion of defective for the 200 rods as Π . We test:

$$H_0 : \Pi \leq 0.1, \quad H_1 : \Pi - 0.1 > \delta = 0.1$$

We test a sample of size $n=20$ and reject H_0 if more than $c=3$ rod fails to meet specifications.

1. What is α ?
2. What is β ?
3. *Can you draw the OC curve?
4. Keep α and n the same, if we want to make $\beta = 0.2$, what H_1 will you set?
5. To make $\alpha = 0.05$, assume $n = 10$, what value of c will you choose?
6. **To make $\alpha = 0.05$, $\beta = 0.3$, what value of n will you choose?