

VE401 RECITATION CLASS NOTE4

Multivariate Random Variables

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1 Definition

Discrete:

Let S be a sample space and Ω a countable subset of \mathbb{R}^n . A discrete multivariate random variable is a map

$$X : S \rightarrow \Omega$$

together with a function $f_X : \Omega \rightarrow \mathbb{R}$ with the properties that

- (i) $f_X(x) \geq 0$ for all $x = (x_1, \dots, x_n) \in \Omega$ and
- (ii) $\sum_{x \in \Omega} f_X(x) = 1$

Continuous: Let S be a sample space. A continuous multivariate random variable is a map

$$X : S \rightarrow \mathbb{R}^n$$

together with a function $f_X : \mathbb{R}^n \rightarrow \mathbb{R}$ with the properties that

- (i) $f_X(x) \geq 0$ for all $x = (x_1, \dots, x_n) \in \mathbb{R}^n$ and
- (ii) $\int_{\mathbb{R}^n} f_X(x) dx = 1$

2 Density and Independence

2.1 Marginal Density

$$f_{X_k}(x_k) = \sum_{x_1, \dots, x_{k-1}, x_{k+1}, \dots, x_n} f_{\mathbf{X}}(x_1, \dots, x_n)$$

$$f_{X_k}(x_k) = \int_{\mathbb{R}^{n-1}} f_X(x) dx_1 \dots dx_{k-1} dx_{k+1} \dots dx_n$$

2.2 Conditional Density

$$f_{X_1|x_2}(x_1) := \frac{f_{X_1 X_2}(x_1, x_2)}{f_{X_2}(x_2)} \quad \text{with } f_{X_2}(x_2) > 0$$

2.3 Independence

Two continuous random variables are independent if:

$$f_X(x_1, x_2) = f_{X_1}(x_1) \cdot f_{X_2}(x_2)$$

A full set of n components of a multivariate random variable is independent if:

$$f_X(x_1, \dots, x_n) = f_{X_1}(x_1) \dots \cdot f_{X_n}(x_n)$$

Question: Visualization

X and Y are continuous random variables. X takes on values between 0 and 2 while Y takes on values between 0 and 1. Their joint PDF is indicated below in the graph.

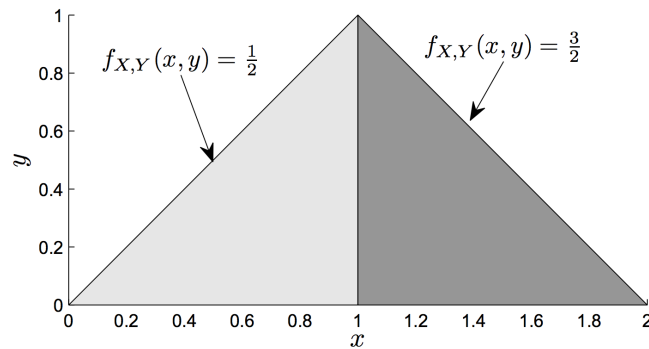


Figure 1: The exponential distribution

1. Are X and Y independent?
2. Find $f_X(x)$ and $f_{Y|X}(y|0.5)$.

Answer: Visualization

1. If X and Y to be independent, any observation of X should not give any information on Y . But in the graph, we see X is observed to be equal to 0, then Y must be 0. You can also describe it mathematically:

$$f_{(X,Y)}(0,0) \neq f_X(0) \cdot f_Y(0)$$

Since $f_{(X,Y)}(0,0) = \frac{1}{2}$, $f_X(0) = 0$, $f_Y(0) = 2$.

You can also think in other ways...

- 2.

$$f_X(x) = \begin{cases} x/2, & \text{if } 0 \leq x \leq 1 \\ -3x/2 + 3, & \text{if } 1 < x \leq 2 \\ 0, & \text{otherwise} \end{cases}$$

$$f_{Y|X}(y|0.5) = \begin{cases} 2, & \text{if } 0 \leq y \leq 1/2 \\ 0, & \text{otherwise} \end{cases}$$

3 Expectation

Definition:

$$E[X] = \begin{pmatrix} E[X_1] \\ \vdots \\ E[X_n] \end{pmatrix}$$

$$E[X_k] = \sum_{x_k} x_k f_{X_k}(x_k) = \sum_{x \in \Omega} x_k f_X(x)$$

$$E[X_k] = \int_{\mathbb{R}} x_k f_{X_k}(x_k) dx_k = \int_{\mathbb{R}^n} x_k f_X(x) dx$$

Property:

For a function $\varphi : \mathbb{R}^n \rightarrow \mathbb{R}$:

$$E[\varphi \circ X] = \sum_{x \in \Omega} \varphi(x) f_X(x), \quad \text{or} \quad E[\varphi \circ X] = \int_{\mathbb{R}^n} \varphi(x) f_X(x) dx$$

4 Variance and Covariance

4.1 Definition

Covariance

$$Cov[X, Y] = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y]$$

which comes from:

$$\begin{aligned}\text{Var}[X + Y] &= \text{E} [((X + Y) - \text{E}[X + Y])^2] \\ &= \text{E} [((X - \text{E}[X]) + (Y - \text{E}[Y]))^2] \\ &= \text{E} [(X - \text{E}[X])^2 + (Y - \text{E}[Y])^2 + 2(X - \text{E}[X])(Y - \text{E}[Y])] \\ &= \text{Var}[X] + \text{Var}[Y] + 2\text{E}[(X - \text{E}[X])(Y - \text{E}[Y])]\end{aligned}$$

Variance

The covariance matrix for a multivariate random variable X is defined as:

$$\text{Var}[X] = \begin{pmatrix} \text{Var}[X_1] & \text{Cov}[X_1, X_2] & \dots & \text{Cov}[X_1, X_n] \\ \text{Cov}[X_1, X_2] & \text{Var}[X_2] & \ddots & \vdots \\ \vdots & \ddots & \ddots & \text{Cov}[X_{n-1}, X_n] \\ \text{Cov}[X_1, X_n] & \dots & \text{Cov}[X_{n-1}, X_n] & \text{Var}[X_n] \end{pmatrix}$$

4.2 Property

1. $\text{Cov}[X, Y] = \text{Cov}[Y, X]$
2. $\text{Cov}[X, X] = \text{Var}[X]$
3. $\text{Var}[CX] = C\text{Var}[X]C^T$, $C \in \text{Mat}(n \times n; \mathbb{R})$ is a constant matrix with real coefficients.

5 *Discussion: Covariance-Linearity-Independence

Covariance and Independence:

1. X and Y are independent $\rightarrow \text{Cov}[X, Y] = 0$.
2. $\text{Cov}[X, Y] = 0 \not\Rightarrow X$ and Y are independent.

Therefore, covariance is not a measure of independence.

Then, what does covariance measure?

Covariance and Linearity:

Covariance measures "linearity". In other words, it shows how much is the relationship between X and Y is like the form " $Y = a + bX$ ".

Let's first have a mathematical taste of the relationship of covariance and linearity. (You do not need to understand exactly clear now.)

We wish to find a best estimation of the linear relationship between variables X and Y , such that the errors between $a + bX$ and Y are minimized. Also notice hopefully we wish $b \neq 0$

to satisfy this is a "linear relationship". The method we choose is to find such a pair of (a, b) making $e = E[(Y - (a + bX))^2]$ minimized. Since e is a function of a and b:

$$e = E[(Y - (a + bX))^2] = E(Y^2) + b^2E(X^2) + a^2 - 2bE(XY) + 2abE(X) - 2aE(Y)$$

To make e minimized, we simply make

$$\begin{cases} \frac{\partial e}{\partial a} = 2a + 2bE(X) - 2E(Y) = 0 \\ \frac{\partial e}{\partial b} = 2bE(X^2) - 2E(XY) + 2aE(X) = 0 \end{cases}$$

Solve the above equations, we obtain:

$$\begin{cases} b = \frac{\text{Cov}[X,Y]}{\text{Var}[X]} \\ a = E(Y) - bE[X] \end{cases}$$

Now we notice, when $\text{Cov}[X, Y] = 0$, $b = 0$; showing X and Y are unlikely to have a linear relationship.

Just relax, let's see a concrete example.

Question: Covariance, Linearity, and Independence

X is a continuous random variable with:

$$f_X(x) = \begin{cases} \frac{1}{2}, & -1 < x < 1 \\ 0, & \text{otherwise} \end{cases}$$

Let $Y = X^2$. Then:

1. What is $\text{Cov}[X, Y]$?
2. Are X and Y independent?
3. Are X and Y linearly related?

Answer: Covariance, Linearity, and Independence

By definition, $Cov[X, Y] = E[XY] - E[X]E[Y] = E[X^3] - E[X]E[X^2]$

And recall, $E[X^k] := \int_{-\infty}^{\infty} x^k \cdot f_X(x)$

Then it's clear $E[X] = E[X^3] = 0$. Hence:

1. $Cov[X, Y] = 0$
2. X and Y are dependent
3. X and Y are not linearly related

6 Pearson Correlation Coefficient

Definition

$$\rho_{XY} = \frac{Cov[X, Y]}{\sqrt{Var[X]Var[Y]}}$$

This idea comes from "the covariance of the standardized X and Y". First standardize X and Y to \tilde{X} and \tilde{Y} , and Then calculate $Cov[\tilde{X}, \tilde{Y}]$.

$$\begin{aligned} Cov[\tilde{X}, \tilde{Y}] &= E[\tilde{X}\tilde{Y}] - E[\tilde{X}]E[\tilde{Y}] \\ &= \frac{Cov[X, Y]}{\sqrt{Var[X]Var[Y]}} \end{aligned}$$

So you can interpret ρ_{XY} as the covariance of the standardized variables, and it is a measure of linearity.

Property

1. $-1 \leq \rho_{XY} \leq 1$
2. $|\rho_{XY}| = 1$ if and only if there exist numbers $\beta_0, \beta_1 \in \mathbb{R}$, $\beta_1 \neq 0$, such that almost surely $Y = \beta_0 + \beta_1 X$.
3. Further $\rho_{XY} = 1$ leads to almost surely $\tilde{X} - \tilde{Y} = 0$, and $\rho_{XY} = -1$ leads to almost surely $\tilde{X} + \tilde{Y} = 0$.

6.1 Fisher Transformation

$Var[\tilde{X} - \tilde{Y}] = 2 - 2\rho_{XY}$, and $Var[\tilde{X} + \tilde{Y}] = 2 + 2\rho_{XY}$. The Fisher transformation is defined as:

$$\ln\left(\sqrt{\frac{Var[\tilde{X} + \tilde{Y}]}{Var[\tilde{X} - \tilde{Y}]}}\right) = \frac{1}{2} \ln\left(\frac{1 + \rho_{XY}}{1 - \rho_{XY}}\right) = \text{Artanh}(\rho_{XY}) \in \mathbb{R}$$

Then we also have:

$$\rho_{XY} = \tanh \left(\ln \left(\frac{\sigma_{\tilde{X}+\tilde{Y}}}{\sigma_{\tilde{X}-\tilde{Y}}} \right) \right)$$

6.2 ρ_{XY} and Linearity

1. The closer $|\rho_{XY}|$ is to 1, the more likely X and Y have a linear relationship.
2. **Positive correlation:** $\rho_{XY} > 0$, $Var[\tilde{X} - \tilde{Y}] < Var[\tilde{X} + \tilde{Y}]$. X and Y tend to be more like $\tilde{X} = \tilde{Y}$, in other words, more positively related. When X is larger, then Y is likely to be larger.
3. **Negative correlation:** $\rho_{XY} < 0$, X and Y tend to be more like $\tilde{X} = -\tilde{Y}$. When X is larger, then Y is likely to be smaller.

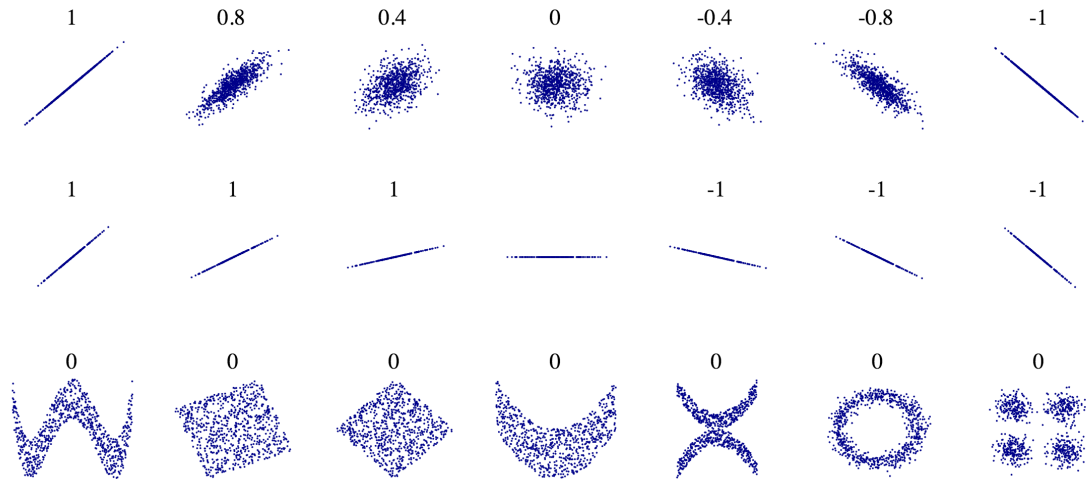


Figure 2: The Pearson Correlation Coefficient

6.3 Bivariate Normal Distribution

X and Y should each follow a normal distribution, but may be not independent. Then the joint density function is:

$$f_{XY}(x, y) = \frac{1}{2\pi\sigma_X\sigma_Y\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)} \left[\left(\frac{x-\mu_X}{\sigma_X} \right)^2 - 2\rho \left(\frac{x-\mu_X}{\sigma_X} \right) \left(\frac{y-\mu_Y}{\sigma_Y} \right) + \left(\frac{y-\mu_Y}{\sigma_Y} \right)^2 \right]}$$

Property:

1. $-1 < \rho < 1$, $\rho = \rho_{XY}$
2. $\mu_X = E[X]$, $\delta_X^2 = Var[X]$. Same for Y.
3. $\rho = 0 \iff$ X and Y are independent. (Notice this is special.)

7 Hypergeometric Distribution

Interpretation:

A total of N balls, r red and $N - r$ black. Draw n balls out without putting back. Assume $r > n$ and $N - r > n$. The random variable X describes the number of red balls in the n drawn balls.

Comment:

It is a sequence of identical but not independent Bernoulli trials. Each draw is a Bernoulli trial with $p = \frac{r}{N}$.

Features:

1. N, r, n are the parameters

$$2. f_X(x) = \frac{\binom{r}{x} \binom{N-r}{n-x}}{\binom{N}{n}}$$

$$3. E[X] = n \frac{r}{N}$$

$$4. \text{Var } X = n \frac{r}{N} \frac{N-r}{N} \frac{N-n}{N-1}$$

Approximate Hypergeometric with Binomial:

Recall for a Binomial distribution, $E[X] = np$, $\text{Var } X = npq$.

If the sampling fraction $\frac{n}{N}$ is small (less than 0.05), we can use $E[X] = n \frac{r}{N} = p$, $n \frac{N-r}{N} = q$ to estimate the Hypergeometric distribution with a Binomial distribution.

We can interpret as: when the sample size n is small, it does not affect the left balls much. So, the trials are approximately independent.

Question: Covariance and Hypergeometric

There are totally N balls, r red and $N-r$ black. We draw 2 balls out one by one without putting back. And we define three Bernoulli random variables representing the below events:

X_1 : Red ball at the first draw.

X_2 : Red ball at the second draw.

X_3 : Black ball at the second draw.

Then, what is $\text{Cov}[X_1, X_2]$, and what is $\text{Cov}[X_1, X_3]$?

And, what is $\rho_{X_1 X_2}$, and what is $\rho_{X_1 X_3}$?

Answer: Covariance and Hypergeometric

We have calculated $\text{Cov}[X_1, X_2]$ in class.

$$\text{Cov}[X_1, X_2] = E[X_1 X_2] - E[X_1]E[X_2] = -\frac{1}{N} \frac{r(N-r)}{N(N-1)} < 0.$$

Similarly since $X_1 X_3$ is still a Bernoulli trial, hence $E[X_1 X_3] = P[X_1 = 1 \text{ and } X_3 = 1] = \frac{r}{N} \frac{N-r}{N-1}$.

$$\text{Then } \text{Cov}[X_1, X_3] = E[X_1 X_3] - E[X_1]E[X_3] = \frac{1}{N} \frac{r(N-r)}{N(N-1)} > 0.$$

Further with $\text{Var}[X_1] = \text{Var}[X_2] = \frac{r}{N}(1 - \frac{r}{N})$ and $\text{Var}[X_3] = \frac{N-r}{N}(1 - \frac{N-r}{N})$. We obtain:

$$\rho_{X_1 X_2} = -\frac{1}{N-1} < 0, \text{ negative correlated.}$$

$$\rho_{X_1 X_3} = \frac{1}{N-1} > 0, \text{ positive correlated.}$$

This also help you understand the positive and negative ρ_{XY} better.

1. When $X_1 = 1$, X_2 is more likely to be 0, and X_3 is more likely to be 1.
2. When N becomes larger, $|\rho_{X_1 X_2}|$ and $|\rho_{X_1 X_3}|$ become smaller. So the effect of X_1 on X_2 and X_3 becomes less. When N is large enough, ρ is close to 0, and we estimate the two trials as independent.